Informal Jobs and Non-fatal Occupational Injuries

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Objectives: In Brazil, workers without a formal job contract represent approximately half of the labor force but there are no official statistics on occupational injuries for them. This study estimates the annual incidence of non-fatal work-related injuries for workers with and without job contracts and examines gender differences.

Methods: This is a community-based study carried out with a random cluster area sample of the residents of Salvador, a city with 2.7 million inhabitants, the capital of the state of Bahia, northeast Brazil. Individuals from 18 to 65 years of age who reported having a paid job comprise the study population (n = 2907). Data were obtained in individual household interviews with questionnaires applied by trained field workers.

Results: The overall estimated annual incidence rate (IR) was 5.6/100 full-time equivalent workers (FTE). The incidence of injuries differed between workers with informal (IR = 6.2/100 FTE) and formal jobs (IR = 5.1/100 FTE), and according to gender (IR = 5.8/100 FTE for female and 5.5/100 FTE for male), but these differences were not statistically significant. Statistically significant positive associations between informal jobs and non-fatal work injuries were observed among women with medium education [incident rate ratio (IRR) 2.02, 95% CI 1.00–4.00] and women with black skin (IRR 1.71, 95% CI 0.99–2.97) who perceived a job as dangerous (IRR 2.00; 95% CI 1.09–3.64) or who had no occupational training (IRR 2.08; 95% CI 1.05–4.20).

Conclusions: This study shows that non-fatal work injuries are a common health problem among adults in urban Brazil, regardless of the type of job contract or gender, which points to a need to improve workers' health and safety programs for formal and informal hired workers.

Keywords: contingent work; informal jobs; occupational injuries; precarious jobs; work accidents; work injuries

INTRODUCTION

Occupational injuries are one of the most important preventable health problems throughout the world. They are potentially fatal or disabling and affect mainly individuals in their productive years causing major economic and social impact. In Brazil, the second largest country in the Western hemisphere, the economic relevance of work accidents is expressed by the annual estimated cost of R\$12.5 billion for company owners and of R\$20 billion for the country (Pastore, 1999), although these data are limited to legal firms and formally hired employees. For the year 1997, the Ministry of Labor and Employment

consequence of work-related injuries and that 11152 individuals received compensation for permanent disability caused by work accidents (Ministério do Trabalho e Emprego, 2000). These statistics are recognized as being of limited quality and restricted to workers having formal job contracts (Wünsch-Filho, 1999). However, informal job contracts are held by half of the Brazilian labor force and by a growing segment of workers in several other developing and developed countries. The growth of this form of employment is described as one of the consequences of changes in the global economy and in the labor force structure (Cacciamali, 1997). According to Quinlan et al. (2001), normal jobs, which means full-time and relatively secure jobs, have been replaced by non-standard or atypical ones, which have received several denominations. In Brazil.

estimated that 2802 workers lost their lives as a

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informal jobs represent those that lack a formal contract and therefore do not officially exist. As a result, informal workers, whether employed or autonomous, are not covered by labor legislation available to formally hired individuals, which ensure regular workplace inspections, occupational hazards control and workers' health monitoring and protection (Santana *et al.*, 1997b), widely recognized as important components of effective policies in reducing occupational diseases and work injuries.

Although not covered by official protection, workers having precarious employment, such as selfemployment, home-based or contingent jobs, are more likely to be in dangerous occupations than formally hired individuals. In a recent review (Quinlan et al., 2001), of 17 studies focusing on the relationship between precarious work arrangements and work injuries, 16 reported higher incidences of fatal, non-fatal or more severe injuries than the referent group. Workers with precarious jobs are also less likely to report injuries or have an awareness of work hazards and they receive less training and supervision (Quinlan et al., 2001). In the unregulated informal sector of the economy, where precarious job contracts prevail, illegal small enterprises are commonly involved in rudimentary operational processes, in which safety is not carefully controlled, which in turn may contribute to an even higher occurrence of work-related injuries and diseases (Loewenson, 1998). Women are more likely to be employed in the informal sector or to have informal job contracts, but research shows that they are less likely to sustain work injuries than men (Biddle and Blanciforti, 1999). However, it is possible that women's increased participation in the labor force and in traditional male jobs may cause changes (Quinlan et al., 2001) that need to be studied.

In this study, based on data from a community-based survey, differences concerning the annual incidence of non-fatal work-related injuries according to the type of job contract and gender are examined in an urban area of Brazil. The incidence distribution is also estimated according to sociodemographic and occupational variables.

MATERIALS AND METHODS

This is a population-based survey carried out in the city of Salvador, capital of the state of Bahia, Brazil. At the time of the study, the year 2000, there were 2.7 million inhabitants, the third largest urban concentration in the country. Its population lives mostly in poor slum areas surrounding upper and middle class neighborhoods and have one of the highest unemployment rates in the country (IBGE, 1999). The study population comprised individuals from 18 to 65 years of age who reported having a paid job and were living in the study area. Sample size was defined

based on the estimated prevalence of informal jobs in the study area and the expected incidence of non-fatal occupational injuries reported in another study (Loewenson, 1998). To identify the study population, a cluster area single stage random sample was designed using 1:12000 and 1:2000 maps provided by a state urban development company. For each sampled area, members of the research staff visited the area and drew detailed hand-made maps showing all local constructions, boundaries and landmarks to be used in undeveloped areas. The number of selected sub-areas was calculated based on the average number of domiciles (86.6) and the expected mean number of family members in the age range of the interest (3.8) (IBGE, 1999). Three of the 32 selected areas were uninhabited, leaving 29, of which local community organizations and the leadership were contacted and asked to participate in the research, and to support operational aspects. The reasons for choosing this sampling design were to overcome difficulties in identifying residential addresses, to ensure field workers' safety and to define environmental and spatial variables. Also, the ethnographic component of this study focused on neighborhood and social support networks.

Trained field workers visited each household, enlisted all family members, registered whether they had a paid job and collected sociodemographic information. All paid workers were individually interviewed, after informed consent was obtained from each participant. Data collection was carried out from July to December 2000, thus avoiding the high tourist season, when informal jobs are more intense and follow patterns distinct from those for the rest of the year. Each household was visited by a study supervisor to check the data collected by field interviewers. When needed, missing information was recorded by telephone. Research instruments were developed following a conceptual map based on relevant content identified in workshops conducted with members of health and labor institutions, non-governmental organizations and the academic community and in questionnaires utilized in previous studies (Barata and Ribeiro, 1998; Fassa, 2000). The language adequacy and operational feasibility of the overall research strategy were tested in a pilot study.

Occupational injuries are defined as 'any damage inflicted on the body by transference of energy during work or when commuting that involves a short duration between exposure and identifiable effects after the event or circumstance' (Hagberg *et al.*, 1997). Musculoskeletal injuries and poisoning were excluded. Perception and recognition of the occupational nature of accidents could be biased by workers with informal jobs, who do not obtain any direct compensation or social benefits out of occupational injuries. To minimize such bias, respondents were asked 'Have you ever had an accident of any type in

your life?', followed by other questions concerning the study reference period, the last 12 months before the interview. When the answer was positive, the respondent was invited to provide a narrative about what happened, how and when, which was recorded as a written narrative. After this, the trained interviewer classified the injury as 'typical' occupational (meaning at the place of work) or when commuting. Severe injuries were those that caused three or more days off work. When an occupational injury was reported, a questionnaire soliciting detailed information was applied. When more than one injury occurred during the referent time, only the most recent was considered. Occupational history was also recorded, considering simultaneous jobs, a common feature in this urban area. A detailed description of the main current job, defined as that involving the longest work-time or highest remuneration, was recorded. For each job, daily work hours and work days per week were also recorded to allow estimation of the total hours of work for each person. Brazilian law requires job contracts to be registered in an official booklet called the Work Card, a widely known symbol of formal placement in the labor market. The type of job contract for participants in the study was measured by whether they had job contracts registered in this Work Card. Other variables of interest were age, sex, skin color (classified as black, including mulattoes, or non-black) and educational level (categorized as: low, illiterates and those who completed junior high education; medium, high school; high, college level). Socio-economic status was defined based on which of nine assets or goods the family owned: car, computer, washing machine, dishwasher, video player, laser disc player, microwave, telephone, beach house. The total number of items was categorized as: low, less than three items; medium, three to five items; high, more than five items. Because of gender differences, trades were coded distinctively for women: services (housemaids, housekeepers, babysitters, cooks and servants), retail, administration/education and other. For men the trades were: service, retail, manufacturing, construction, transport/safety and other. Other occupational variables were type of workplace (whether a firm or plant, street or domicile), shift-work/night work, perception of the job as dangerous, occupational training and sub-contracting, all coded as yes/no.

The annual incidence rate was estimated by dividing the total number of reported non-fatal work injuries by the summation of person-time, calculated in hours of work in the last 12 months, taking 100 full-time worker equivalents (FTE) as 40 h/week in a year of 50 weeks or 200000 h of work in a year (Mital *et al.*, 1999). Although the study design is cross-sectional in nature, accidents are of short duration and it is appropriate to consider them as incident

events even when based on self-reporting. Statistical inference for differences in proportions was by Pearson χ^2 test. Crude incidence ratios were estimated using Poisson regression models and confidence intervals based on the Wald test (Stockes et al., 2000). Double data entry allowed a complete check of the database and corrections when needed. Data analysis was performed with SAS 8.11 (SAS Inc., 1998), adjusting the final results for sampling design using PROC SURVEYMEANS. The research protocol was reviewed and approved by the Internal Review Board of the Hospital Prof. Edgard Santos, Federal University of Bahia and the Center for Health Research, University of Texas at Houston. The research staff were informed of all ethical issues involved and their related responsibilities.

RESULTS

A total of 3528 paid workers aged 18-65 vr were identified in the original study. From them, 212 (6.0%) refused to participate, 43 (1.2%) had retired, 260 (7.4%) were public officers under special labor legislation, 48 (1.4%) were entrepreneurs and 31 (0.9%) were autonomous professionals. All of these workers, as well as seven with missing labor market placement data, were excluded. The remaining study population comprised 2907 individuals, most of whom were male (1537, 52.9%) and did not have formal job contracts (1517, 52.2%). Only one case reported two occupational injuries in the study period. The estimated overall annual incidence rate of work accidents was 5.6/100 FTE. The incidence of injuries differed between estimates for workers with informal [incidence rate (IR) = 6.2/100 FTE] and formal jobs (IR = 5.1/100 FTE) and according to gender (IR = 5.8/100 FTE for female and 5.5/100FTE for male), but the differences were not statistically significant. Among women, the incidence rate of at-work injuries was estimated as 5.9/100 FTE in the informal sector and 3.4/100 FTE in the formal sector, a statistically significant difference [IR = 1.75, 95%] confidence interval (CI) 1.02-2.99]. No differences were found for at-work injuries in the male group, or for commuting injuries across gender or type of job contract.

Women reporting informal job contracts were younger, had less education and were more likely to be of low socio-economic status than those who reported having formal contracts. They were also more likely to work on the streets or in domiciles, be a housemaid or housekeeper or have a job in the retail trade, a shift-work schedule and no occupational training (Table 1). There were no differences concerning skin color or perception of the job as dangerous. Men with informal jobs showed similar sociodemographic and occupational profiles, except that socio-economic status did not differ according to

Table 1. Characteristics of the study population according to type of job contract and gender

| Variables | Type of job contract | | | | | | |
|---|----------------------|------|----------|------|----------|--------|--|
| | Informal | | Formal | | Total | | |
| | n = 1517 | % | n = 1390 | % | n = 2907 | 100.0% | |
| Female | | | | | | | |
| Age under 22 years old ^a | 101 | 13.1 | 52 | 8.7 | 153 | 11.2 | |
| Black skin color | 484 | 62.7 | 365 | 61.0 | 849 | 62.0 | |
| Less than elementary education ^b | 406 | 52.6 | 205 | 34.3 | 611 | 44.6 | |
| Low SES ^b | 425 | 55.1 | 271 | 45.5 | 696 | 50.9 | |
| Trade ^b | | | | | | | |
| Services/ housemaids | 248 | 32.1 | 166 | 27.8 | 414 | 30.2 | |
| Retail | 211 | 27.3 | 129 | 21.6 | 340 | 24.8 | |
| Administration/education | 103 | 13.3 | 144 | 24.1 | 247 | 18.0 | |
| Other ^b | 210 | 27.2 | 159 | 26.6 | 369 | 26.9 | |
| Workplace ^b | | | | | | | |
| Firm or plant | 189 | 24.5 | 467 | 78.1 | 656 | 47.9 | |
| Street | 94 | 12.2 | 3 | 0.5 | 97 | 7.1 | |
| Domicile | 276 | 35.8 | 23 | 3.9 | 299 | 21.8 | |
| Other | 213 | 27.6 | 105 | 17.6 | 318 | 23.2 | |
| Shift-work/night workb | 160 | 26.0 | 120 | 22.7 | 280 | 24.5 | |
| Perceive job as dangerous | 206 | 26.8 | 167 | 28.0 | 373 | 27.3 | |
| Occupational training ^b | 258 | 33.6 | 323 | 54.4 | 581 | 42.7 | |
| Subcontracting | 11 | 4.4 | 30 | 5.7 | 41 | 5.3 | |
| Male | | | | | | | |
| Age under 22 years old ^b | 89 | 12.0 | 64 | 8.1 | 153 | 10.0 | |
| Black skin color | 460 | 61.7 | 474 | 59.9 | 934 | 60.8 | |
| Less than elementary education ^b | 452 | 60.7 | 355 | 44.8 | 807 | 52.5 | |
| Low SES | 404 | 54.5 | 398 | 50.3 | 802 | 52.3 | |
| Trade ^b | | | | | | | |
| Services | 38 | 5.1 | 77 | 9.7 | 115 | 7.5 | |
| Retail | 191 | 25.6 | 163 | 20.6 | 354 | 23.0 | |
| Industry | 37 | 5.0 | 86 | 10.9 | 123 | 8.0 | |
| Construction | 181 | 24.3 | 95 | 12.0 | 276 | 18.0 | |
| Transport/safety | 78 | 10.5 | 208 | 26.2 | 286 | 18.6 | |
| Other | 220 | 29.5 | 163 | 20.6 | 383 | 24.9 | |
| Workplace ^b | | | | | | | |
| Firm or plant | 250 | 33.5 | 661 | 83.4 | 911 | 59.3 | |
| Street | 183 | 24.6 | 75 | 9.5 | 258 | 16.8 | |
| Domicile | 217 | 29.1 | 14 | 1.8 | 231 | 15.0 | |
| Other | 95 | 12.8 | 42 | 5.3 | 137 | 8.9 | |
| Shift-work/night workb | 127 | 17.3 | 234 | 29.6 | 361 | 23.7 | |
| Perceive job as dangerous ^a | 384 | 51.7 | 428 | 54.6 | 812 | 53.2 | |
| Occupational training ^b | 349 | 47.4 | 530 | 66.9 | 879 | 57.5 | |
| Subcontracting ^c | 25 | 12.0 | 44 | 5.6 | 69 | 7.0 | |

^aPearson $\chi^2 P < 0.05$.

type of job contract (Table 1). Men worked primarily in trades where informal jobs prevail, notably construction and retail.

The annual incidence rate of non-fatal work injuries among women was higher in the informal job

group (7.1/100 FTE) than among women with formal contracts (4.6/100 FTE), although not statistically significant (Table 2). Notable positive associations between informal jobs and non-fatal work injuries were observed among women with medium educa-

bPearson $\chi^2 P < 0.001$.

^cPearson $\chi^2 P < 0.01$.

Table 2. Crude annual incidence rate of non-fatal work accidents per 200 000 hours of work a year, according to sociodemographic variables, by type of job contract and gender

| Variable | Type of job | contract | Crude incidence ratio ^a | 95% confidence interval | | |
|-----------------------|-----------------------|--|--|----------------------------|--------|-------------|
| | Informal | | | | Formal | |
| | Person-h ^b | Incidence rate/100 FTE ^c | Person-h | Incidence rate/100 FTE | iulio | |
| Female ($n = 1370$) | 123240 | 7.1 | 22518 | 4.6 | 1.54 | 0.97-2.51 |
| Age (yr) | | | | | | |
| 18-21 | 16255 | 6.2 | 10792 | 3.6 | 1.72 | 0.33-8.83 |
| 22-40 | 70564 | 7.1 | 76323 | 4.2 | 1.69 | 0.90-3.16 |
| 40-65 | 37421 | 7.5 | 35403 | 5.6 | 1.34 | 0.60-3.02 |
| Skin color | | | | | | |
| Black | 75834 | 9.0 | 758 960 | 5.3 | 1.70 | 0.99-2.97 |
| Non-black | 48406 | 4.2 | 467558 | 3.4 | 1.21 | 0.48-3.08 |
| Education | | | | | | |
| Elementary or less | 66496 | 6.9 | 447475 | 5.8 | 1.18 | 0.60-2.33 |
| High school | 48836 | 8.3 | 686493 | 4.1 | 2.02 | 1.00-4.00 |
| College | 8908 | 2.5 | 92550 | 2.2 | 1.14 | 0.07-18.30 |
| Socio-economic status | | | | | | |
| Low | 67955 | 7.5 | 595703 | 4.7 | 1.59 | 0.82-3.05 |
| Medium | 347655 | 6.9 | 437 575 | 5.0 | 1.37 | 0.60-3.11 |
| High | 216630 | 6.5 | 189640 | 3.2 | 2.04 | 0.53-7.90 |
| Male $(n = 1537)$ | 1525113 | 5.5 | 1690210 | 5.4 | 1.01 | 0.67 - 1.54 |
| Age (yr) | | | | | | |
| 18-21 | 141560 | 5.7 | 130678 | 6.1 | 0.92 | 0.23-3.69 |
| 22-40 | 758324 | 6.9 | 1129054 | 5.5 | 1.24 | 0.74-2.10 |
| 40-65 | 625229 | 3.8 | 430478 | 5.1 | 0.76 | 0.33-1.70 |
| Skin color | | | | | | |
| Black | 942315 | 5.7 | 999626 | 6.6 | 0.86 | 0.52-1.44 |
| Non-black | 582798 | 5.1 | 690584 | 3.8 | 1.36 | 0.65-2.87 |
| Education | | | | | | |
| Elementary or less | 905964 | 7.1 | 768310 | 7.5 | 0.93 | 0.57-1.55 |
| High school | 518829 | 3.5 | 816936 | 3.4 | 1.01 | 0.44-2.34 |
| College | 100320 | 2.0 | 104964 | 5.7 | 0.35 | 0.04-3.35 |
| Socio-economic status | | | | | | |
| Low | 758025 | 6.1 | 858168 | 7.0 | 0.87 | 0.50-1.49 |
| Medium | 477943 | 5.0 | 572150 | 4.9 | 1.02 | 0.47-2.22 |
| High | 280733 | 5.0 | 259892 | 1.5 | 3.20 | 0.67 -15.60 |

^aBased on Poisson regression; confidence intervals estimated with the Wald method.

tion [incidence rate ratio (IRR) = 2.02, 95% CI 1.00–4.00] and women with black skin (IRR = 1.71, 95% CI 0.99–2.97). No other statistically significant difference was observed according to the type of job contract for any sociodemographic variable among men.

Regarding occupational variables (Table 3), having an informal job was associated with non-fatal work injuries among women who did not perceive their jobs as dangerous (IRR = 2.00, 95% CI 1.09–3.64) or reported no occupational training (IRR = 2.08, 95% CI 1.05–4.20). There were no statistically significant

differences for occupational variables among men. Elevated incidence ratios for non-fatal injuries were observed in the retail trade for women (IRR = 2.93, 95% CI 0.98–8.91) and men (IRR = 2.00, 95% CI 0.71–5.72) and for women in the housekeeping and cleaning trades (IRR = 1.87, 95% CI 0.87–4.00). None of the associations of non-fatal injuries with informal jobs were statistically significant in any of the trades under analysis among women or men, however (Table 4). When only severe non-fatal work injuries (those resulting in missing work for three or more days) were considered in the analysis, similarly

^bTotals differ because of missing data.

[°]FTE, full-time equivalent workers a year (40 h/week, 50 weeks/yr).

Table 3. Crude annual incidence of non-fatal work accidents per 100 FTE according to occupational variables, by type of job contract and gender

| Variables | Type of job contract | | | | Crude | 95% |
|---------------------------|--------------------------|------------------------------------|-------------|-----------------------|---------------------------------|------------------------|
| | Informal | | Formal | | incidence ratio ^a | confidence interval |
| | Person-h/yr ^b | Incidence /100 FTE ^c | Person-h/yr | Incidence /100 FTE | - | |
| Female (<i>n</i> = 1370) | 1235240 | 7.1 | 1226518 | 4.6 | 1.54 | 0.97-2.51 |
| Workplace | | | | | | |
| Firm or plant | 332260 | 5.4 | 913640 | 4.4 | 1.22 | 0.94 - 2.72 |
| Street | 137946 | 10.1 | 5760 | | | |
| Domicile | 400008 | 6.5 | 60060 | 3.3 | 1.96 | 0.26-14.9 |
| Other | 365026 | 8.2 | 247058 | 5.7 | 1.43 | 0.59 - 3.56 |
| Shift-work/night work | | | | | | |
| Yes | 186096 | 7.5 | 222370 | 6.3 | 1.19 | 0.42 - 3.41 |
| No | 776940 | 6.2 | 830196 | 3.9 | 1.59 | 0.31 - 3.00 |
| Perceive job as dangerous | | | | | | |
| Yes | 368231 | 7.1 | 337297 | 7.1 | 1.00 | 0.45 - 2.17 |
| No | 863 509 | 7.2 | 886197 | 3.6 | 2.00 | 1.09-3.64 |
| Occupational training | | | | | | |
| Yes | 425323 | 5.6 | 620511 | 5.5 | 1.02 | 0.49-2.16 |
| No | 804997 | 7.7 | 596887 | 3.7 | 2.08 | 1.05-4.20 |
| Subcontracting | | | | | | |
| Yes | 14312 | | 51705 | 11.6 | | |
| No | 379424 | 6.3 | 991645 | 4.0 | 1.57 | 0.77 - 3.21 |
| Male $(n = 1537)$ | 1525113 | 5.5 | 1690210 | 5.4 | 1.01 | 0.67 - 1.54 |
| Workplace | | | | | | |
| Firm or plant | 528027 | 2.7 | 1411982 | 5.4 | 0.50 | 0.22-1.10 |
| Street | 399984 | 6.0 | 66596 | 7.2 | 0.83 | 0.31-2.22 |
| Domicile | 429254 | 7.5 | 7908 | 7.2 | 1.04 | 0.14 - 7.84 |
| Other | 167848 | 8.3 | 83724 | 2.4 | 3.49 | 0.42-28.38 |
| Shift-work/night work | | | | | | |
| Yes | 226231 | 7.1 | 500850 | 4.0 | 1.77 | 0.69-4.49 |
| No | 1281014 | 5.3 | 1187632 | 6.1 | 0.87 | 0.54-1.40 |
| Perceive job as dangerous | | | | | | |
| Yes | 820858 | 6.6 | 932409 | 6.9 | 0.95 | 0.56-1.58 |
| No | 699839 | 4.3 | 752785 | 3.7 | 1.15 | 0.56-2.42 |
| Occupational training | | | | | | |
| Yes | 714681 | 5.9 | 1108454 | 5.2 | 1.12 | 0.64-1.97 |
| No | 790356 | 4.8 | 581756 | 5.8 | 0.83 | 0.43 - 1.58 |
| Subcontracting | | | | | | |
| Yes | 45981 | 8.7 | 81643 | 12.2 | 0.71 | 0.38 - 8.52 |
| No | 344014 | 4.6 | 1593111 | 5.1 | 0.90 | 0.71 - 1.94 |

^aBased on Poisson regression; confidence intervals estimated with the Wald method.

no statistically significant positive associations were observed (Table 5).

The distribution of non-fatal work injury cases according to occupation and circumstance or cause of the injury are presented in Table 6. In the retail trade, for both genders, injured workers were predominantly salespeople, with a variety of causes of injury.

Among women, falls and handling cutting tools appeared to be the most common causes of work injuries in the group of informal or formal jobs. Contact with hot substances or surfaces were more common among women having informal jobs in the services or retail trade. Among men, most work injuries involved vehicles; the causes of injury included being struck

^bTotals differ because of missing data.

cFTE, full-time equivalent workers a year (40 h/week, 50 weeks/yr).

Table 4. Crude annual incidence ratios and their respective 95% confidence intervals for type of job contract and non-fatal work accidents according to trade

| Variable | Type of job | contract | Crude | 95% confidence interval | | |
|-----------------------|-------------|-----------------------------------|-------------|-------------------------------|----------------------|-------------|
| | Informal | Informal | | | Formal | |
| | Person-h/yr | Incidence/100 FTE ^b | Person-h/yr | Incidence/100 FTE | _ ratio ^a | mici vai |
| Female ($n = 1370$) | 1235240 | 7.1 | 1226518 | 4.6 | 1.54 | 0.97-2.51 |
| Trade | | | | | | |
| Services/housekeeping | 392857 | 9.7 | 388843 | 5.2 | 1.87 | 0.87-4.00 |
| Retail | 352615 | 8.5 | 278077 | 2.9 | 2.93 | 0.98-8.91 |
| Business/education | 148304 | 5.4 | 261122 | 3.1 | 1.74 | 0.44-7.04 |
| Other | 341464 | 3.5 | 298476 | 6.7 | 0.52 | 0.19 –1.44 |
| Male $(n = 1537)$ | 1525113 | 5.5 | 1690210 | 5.4 | 1.02 | 0.67-1.54 |
| Trade | | | | | | |
| Services/housekeeping | 84932 | 2.5 | 158298 | 12.6 | 0.19 | 0.02-1.46 |
| Construction | 354077 | 7.9 | 199497 | 9.0 | 0.77 | 0.32 - 1.95 |
| Manufacturing | 83828 | 7.2 | 180201 | 9.9 | 0.72 | 0.19-2.65 |
| Retail | 430156 | 5.6 | 361314 | 2.8 | 2.00 | 0.71 - 5.72 |
| Transport/safety | 185124 | 2.2 | 468475 | 3.4 | 0.64 | 0.13-2.98 |
| Other | 394784 | 6.6 | 324045 | 3.7 | 1.78 | 0.68-0.43 |

^aBased on Poisson regression; confidence intervals estimated with the Wald method.

Table 5. Crude annual incidence ratios and their respective 95% confidence intervals for type of job contract and severe (three or more days off work) non-fatal work accidents according to trade

| Variable | Type of job contract | | | | | | | |
|---------------------------|----------------------|-----------------------------------|-------------|----------------------|---------------------------------|------------------------|--|--|
| | Informal | | Formal | | Crude | 95% | | |
| | Person-h/yr | Incidence/100 FTE ^b | Person-h/yr | Incidence/100 FTE | incidence ratio ^a | confidence interval | | |
| Female (<i>n</i> = 1370) | 1235240 | 1.9 | 1226518 | 2.4 | 0.79 | 0.37-1.70 | | |
| Trade | | | | | | | | |
| Services/housekeeping | 392857 | 2.0 | 388843 | 2.6 | 0.77 | 0.21-2.95 | | |
| Retail | 352615 | 3.4 | 278077 | 2.2 | 1.55 | 0.39-6.31 | | |
| Business/education | 148304 | | 261122 | 2.3 | | | | |
| Other | 341464 | 1.2 | 298476 | 2.7 | 0.44 | 0.08 - 2.38 | | |
| Male $(n = 1537)$ | 1525113 | 2.4 | 1690210 | 2.7 | 0.89 | 0.47 - 1.61 | | |
| Trade | | | | | | | | |
| Services/housekeeping | 84932 | 2.5 | 158298 | 5.1 | 0.49 | 0.05-4.17 | | |
| Construction | 354077 | 2.8 | 199497 | 6.0 | 0.35 | 0.12 - 1.70 | | |
| Manufacturing | 83828 | 4.8 | 180201 | 7.9 | 0.61 | 0.20-5.87 | | |
| Retail | 430156 | 1.9 | 361314 | 1.7 | 1.11 | 0.39-1.75 | | |
| Transport/safety | 185 124 | 2.2 | 468475 | 1.7 | 1.29 | 0.23-6.91 | | |
| Other | 394784 | 2.5 | 324045 | 1.9 | 1.32 | 0.33-5.72 | | |

^aBased on Poisson regression; confidence intervals estimated with the Wald method.

by a vehicle, falling from a vehicle and crashes, both while at work and when commuting.

DISCUSSION

This study shows that non-fatal work injuries are a common health problem among adults in urban

Brazil, regardless of the type of job contract or gender. The incidence of non-fatal work injuries was estimated as 5.8/100 FTE/yr, indicating the quantitative importance of occupational injuries, all of which are preventable, in principle, by safety, hygiene and public health measures. The incidence of injury was higher among women with informal jobs compared

^bFTE, full-time equivalent workers a year (40 h/week, 50 weeks/yr).

^bFTE, full-time equivalent workers a year (40 h/week, 50 weeks/yr).

Table 6. Occupations and circumstances/causes^a of non-fatal work accidents according to type of job contract and gender

| Gender/occupation | Circumstances/cause | | | |
|---------------------------------|--|---|--|--|
| | Informal job contract | Formal job contract | | |
| Female | | | | |
| Services | | | | |
| Housemaid/housekeeper | Falls (7), cutting tools (3), contact with hot substances (5) contact with cold substances (2) | Falls (4), cutting tools (2), machine transportation (1) contact with hot surfaces (1), squeezed by a vehicle (1) | | |
| Cook | | Fall (1) | | |
| Cleaners | Fall (1) | | | |
| Other | | Fall (1) | | |
| Retail | | | | |
| Salesperson/vendor | Falls (4), struck by a moving vehicle (1), cutting tools (1), contact with hot substances (2), overextension (1) | Fall (1) | | |
| Cashier | | Falls (2) | | |
| Administrative assistant | Contact with hot substances (1) | Fall (1) | | |
| Waiter | Contact with hot substances (1) | | | |
| Business/education | | | | |
| Teacher | Cutting tools (1), contact with hot substances (2) | Falls (2), cutting tools (1) | | |
| Communication operator | | Fall (1) | | |
| Phone operator | Car crash (1) | | | |
| Other | | | | |
| Industrial operator | | | | |
| Laboratory technician | | Cutting tools (1) | | |
| Nurse attendants | | Cutting tools (1), contact with hot substance (1) | | |
| Hairdresser | Cutting tools (1) | | | |
| Dressmaker | Cutting tools (2), machine handling (1) | | | |
| Dispatcher | | | | |
| Handler | | Fall (1) | | |
| Male | | | | |
| Services | | | | |
| Janitor | | Fall (4), cutting tools (1), overexertion (1) | | |
| Gardener | | Fall (2) | | |
| Construction | | | | |
| Bricklayer/bricklayer assistant | Fall (2), fall from a vehicle (2), cutting tools (1) | | | |
| Scaffolders | | Fall (1) | | |
| Electrician | Gas leaking (1) | Fall (2), squeezed by a machine (1) | | |
| Painter | Struck by a moving vehicle (1) | Chemical substance (1) | | |
| Carpenter | | Struck by a moving vehicle (1) | | |
| Plumbers | Cutting tools (1) | | | |
| Foundry worker | Cutting tools (1) | | | |
| Overall construction | | Fall (1), gunshot (1), car collision (1) | | |
| Manufacturing | | | | |
| Welder | | Struck by a moving vehicle (1) | | |
| Sheet metal worker | Cutting tools (2) | Struck by a moving vehicle (1) | | |
| Operator/maintenance | Machine transportation (1) | Struck by a moving vehicle (3), machine handling (2) | | |
| Electronic operator | Machine handling (1) | | | |
| Other | Cutting tools (3) | | | |

Table 6. Continued

| Gender/occupation | Circumstances/cause | | | | | |
|--------------------------|---|---|--|--|--|--|
| | Informal job contract | Formal job contract | | | | |
| Retail | | | | | | |
| Salesperson | Fall from a vehicle (1), struck by a moving vehicle (1), cutting tools (2), contact with hot surfaces (2) | Cutting tools (1), overextension (1) | | | | |
| Cashier | | Fall (1) | | | | |
| Waiter | Struck by a moving vehicle (1), contact with chemical substance (1) | | | | | |
| Supervisor | Fall from a vehicle (1) | | | | | |
| Stock handler | | Cutting tools (1) | | | | |
| Administrative assistant | | Fall (1) | | | | |
| Telephone operator | Fall (1) | | | | | |
| Transport | | | | | | |
| Dispatcher | Squeezed by machinery (1) | | | | | |
| Motorcycle courier | Struck by a moving vehicle (1) | Struck by a moving vehicle (2) | | | | |
| Other | | | | | | |
| Hairdresser | Fall (1) | | | | | |
| Priest | Fall (1) | | | | | |
| Teacher | | Fall from a vehicle (1) | | | | |
| Communications operator | | Cutting tools (1) | | | | |
| Butcher | | Cutting tools (1) | | | | |
| Driver | | Contact with hot substance (1), overextension (1) | | | | |
| Security guard | | Chemical substance (1) | | | | |
| Print operator | | Machine handling (1) | | | | |
| Artist | Fall from a vehicle (1), cutting tools (1) | | | | | |
| Artisan | Cutting tools (1) | | | | | |

^aData on circumstance/cause for 40 individuals were missing.

with those with formal jobs, although the difference was not statistically significant. In addition, women with informal job contracts were more likely to report work accidents than those having formal contracts when they had a medium level education (high school), did not perceive their job as dangerous or had no job training. No substantial differences between formal and informal workers were found among men, however. Although few statistically significant differences in injury risk between formal and informal workers were observed, the numbers of workers in many groups were small, limiting the power of the study.

The measure of injury occurrence utilized in this study, the annual incidence of non-fatal work injuries per 100 full-time workers a year requires an estimate of hours of work, which is typically not recorded and can be difficult to assess in research. The scarcity of studies using such person-time data limits comparison with other findings. These results are close to the estimated cumulative incidence of 5.0% a year reported by Frumkim and Câmara (1991) for the entire population of Brazil, and a little higher than the estimate of 4.24% for the same age group in the state of São

Paulo in a previous study using a similar methodology (Barata and Ribeiro, 2000). Manufacturing workers are widely known to have an increased risk of work accidents relative to other trades, but the estimates in this study are similar to the 5.6/100 person-yr reported for workers in a steel plant in Minas Gerais (Schoemaker and Barreto, 2000).

Comparisons with other countries are also limited, because surveys looking at non-fatal injuries are rare. Much research is based on secondary analysis of databases created for different purposes and of different natures, coverage and quality. Nevertheless, data from emergency departments in US hospitals indicated an incidence of 2.9/100 FTE/yr for individuals over 15 years of age, corresponding to 3.4/100 FTE/yr among men and 2.1/100 FTE/yr for women (CDC, 1998). The rates reported in this study are lower than those reported in a study from Costa Rica (Wesseling et al., 2001), but that study included pesticide poisoning and the population was rural, where work accidents are known to be more frequent. In a study carried out in Zimbabwe, the estimated frequency of work accidents was higher, 13.1%, but the study population was both rural and urban and the social and economic contexts are quite different from urban Brazil (Loewenson, 1998).

In contrast to the majority of the results reported in a review by Quinlan et al. (2001) on the health consequences of precarious work arrangements, no statistically significant association was found between informal jobs, a precarious form of labor market participation, and non-fatal work accidents, for women or men. Similar results were reported in other surveys conducted in the state of São Paulo (Barata and Ribeiro et al., 2000) and in the metropolitan area of Aracaju (Dantas and Santana, 2002), both in Brazil. This negative finding may be a result of real similarities in work conditions in the informal and formal sectors in Brazil, where poor enforcement of labor legislation is widespread. Formally hired workers may also be holding precarious jobs, a possibility that points to the need to look at an improvement in work conditions for formal as well as informal workers. The weak influence of the type of labor contract on injury incidence could also be a result of increased under-reporting of work accidents among informal workers compared with formal workers, which may exist because of the compensatory perspective on work accidents among formal workers. In a Finnish study, workers having contingent employment, another form of precarious job status, showed better self-rated health and a lower sickness absence rate than permanent workers (Virtanen et al., 2001), which was explained as the result of differences in the threshold for taking sick leave or working even when sick, a behavior likely to occur among workers who fear losing their jobs. This phenomenon may also exist in Brazil, where unemployment is high and competition for jobs in both the formal and informal sectors is substantial.

It is obvious that the ability to consider some occupational injury risk factors related to work processes, technologies, design of the work environment and safety level were limited in this study because of its exploratory nature and the wide range of job titles and trades involved. Nevertheless, sociodemographic, family and occupational characteristics may give clues that underscore important macrosocial aspects that are relevant to prevention. This study showed the importance of occupational training and risk awareness for the association of informal job contracts and work accidents. Nevertheless, distinct from other studies (Loewenson, 1998; Barata and Ribeiro et al., 2000), younger age was not associated with increased risk of occupational injuries according to gender or type of job contract.

The failure to observe statistically significant associations of injury risk with informal work among men may be a consequence of the increased job-related risks thought to be more common in male-dominated occupations. Greater under-reporting of work injuries may also occur among men, relative to women, who

are often reported as having an increased perception and promptness to speak out about health problems, compared with men (Santana et al., 1997a). From this perspective, it is possible that studies that show men at higher risk of occupational accidents are indeed affected by selection bias because women are more likely to have informal job contracts. When informal workers were included in the study population, gender differences tend to disappear. These results need to be looked at with caution considering the previously mentioned tendency of women to be better informants of health-related events. It is worth noting that, in this study, detailed narratives were registered for each injured worker, as well as information about medical treatment received, procedures that should reduce overestimation of health events.

In occupational epidemiology, general population studies are traditionally conducted with secondary data obtained from official records, despite concerns about poor quality and limited coverage (Leigh et al., 1997; Mital et al., 1999; Wünsch-Filho, 1999). These difficulties are greater in regions where unregistered jobs account for half of the labor force, as in Brazil, and there is no information system that allows access to appropriate occupation-specific population denominators. In addition, health problems of occupational origin are not recognized or adequately registered in health services or work places (Santana et al., 2002). In Brazil, it is possible that with the recent integration of workers' health actions into the scope of the Healthy Family Program, a new strategy of prevention and health promotion that covers informal workers will be implemented (Ministério da Saúde, 2001).

The purpose of this study was exploratory and limited to raising hypotheses. The community-based cross-sectional design is not traditional in occupational injury research, but in the Brazilian context it is the only feasible strategy to obtain reliable estimates of this important public health problem. Informal workers are largely neglected in both official statistics and traditional, industry-based epidemiologic studies. Population-based studies on workers' health are also affected by the dispersion of individuals among a large number of occupations, which tends to limit power. Nevertheless, these studies have the advantage that they represent an existing epidemiological profile of the workforce, which allows the identification of occupational groups that can be studied in more detail with other methodologies. Studies using secondary data and capture-recapture methods or focusing on health care services, such as hospital emergency departments, may help to draw a more complete picture of work injuries in scenarios where informal jobs are common and form a basis to develop well-tailored surveillance models, such as the strategy of sentinel areas (Santana, 2002).

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